

# Example

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[Carroll, *J Med Entomol* **38**:114–117, 2001]

Place tick on clay island surrounded by water, with two capillary tubes: one treated with deer-gland-substance; one untreated.

Does the tick go to the treated or the untreated tube?

| Tick sex | Leg  | Deer sex | treated | untreated |
|----------|------|----------|---------|-----------|
| male     | fore | female   | 24      | 5         |
| female   | fore | female   | 18      | 5         |
| male     | fore | male     | 23      | 4         |
| female   | fore | male     | 20      | 4         |
| male     | hind | female   | 17      | 8         |
| female   | hind | female   | 25      | 3         |
| male     | hind | male     | 21      | 6         |
| female   | hind | male     | 25      | 2         |

Is the tick more likely to go to the treated tube?

## Test for a proportion

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Suppose  $X \sim \text{binomial}(n, p)$ .

Test  $H_0 : p = \frac{1}{2}$  vs  $H_a : p \neq \frac{1}{2}$

Reject  $H_0$  if  $X \geq H$  or  $X \leq L$

Choose  $H$  and  $L$  such that

$$\Pr(X \geq H \mid p = \frac{1}{2}) \leq \alpha/2 \text{ and } \Pr(X \leq L \mid p = \frac{1}{2}) \leq \alpha/2$$

Thus  $\Pr(\text{Reject } H_0 \mid H_0 \text{ is true}) \leq \alpha$ .

**The difficulty:** The binomial distribution is hard to work with. Because of its discrete nature, you can't get **exactly** your desired significance level ( $\alpha$ ).

# Rejection region

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Consider  $X \sim \text{binomial}(n=29, p)$

Test of  $H_0 : p = \frac{1}{2}$  vs  $H_a : p \neq \frac{1}{2}$  at significance level  $\alpha = 0.05$

Lower critical value:

$$q_{\text{binom}}(0.025, 29, 0.5) = 9$$

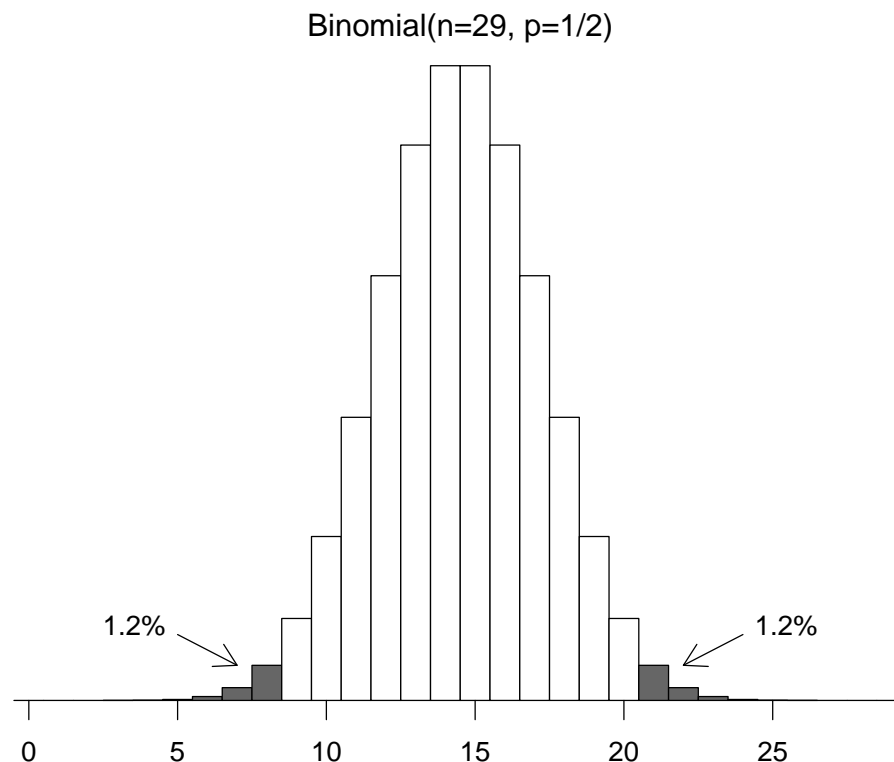
$$\Pr(X \leq 9) = p_{\text{binom}}(9, 29, 0.5) = 0.031 \rightarrow \mathbf{L = 8}$$

Upper critical value:

$$q_{\text{binom}}(0.975, 29, 0.5) = 20$$

$$\Pr(X \geq 20) = 1 - p_{\text{binom}}(20, 29, 0.5) = 0.031 \rightarrow \mathbf{H = 21}$$

Reject  $H_0$  if  $X \leq 8$  or  $X \geq 21$ . (For testing  $H_0 : p = \frac{1}{2}$ ,  $H = n - L$ .)



# Significance level

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Consider  $X \sim \text{binomial}(n=29, p)$

Test of  $H_0 : p = \frac{1}{2}$  vs  $H_a : p \neq \frac{1}{2}$  at significance level  $\alpha = 0.05$

Reject  $H_0$  if  $X \leq 8$  or  $X \geq 21$ .

Actual significance level:

$$\begin{aligned}\alpha &= \Pr(X \leq 8 \text{ or } X \geq 21 \mid p = \frac{1}{2}) \\ &= \Pr(X \leq 8 \mid p = \frac{1}{2}) + [1 - \Pr(X \leq 20 \mid p = \frac{1}{2})] \\ &= \text{pbinom}(8, 29, 0.5) + 1 - \text{pbinom}(20, 29, 0.5) \\ &= 0.024\end{aligned}$$

If we used, instead, “Reject  $H_0$  if  $X \leq 9$  or  $X \leq 20$ ,” the significance level would be:

$$\text{pbinom}(9, 29, 0.5) + 1 - \text{pbinom}(19, 29, 0.5) = 0.061$$

## Example

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Observe  $X = 24$  (for  $n = 29$ )

Reject  $H_0 : p = \frac{1}{2}$  if  $X \leq 8$  or  $X \geq 21$ .

Thus we reject  $H_0$  and conclude that the ticks were more likely to go to the deer-gland-substance-treated tube.

$$\begin{aligned}\text{P-value} &= 2 \times \Pr(X \geq 24 \mid p = \frac{1}{2}) \\ &= 2 * (1 - \text{pbinom}(23, 29, 0.5)) \\ &= 5/10,000.\end{aligned}$$

Alternatively: `binom.test(24, 29)`

## Example 2

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Observe  $X = 17$  (for  $n = 25$ ); assume  $X \sim \text{binomial}(n=25, p)$

Test  $H_0 : p = \frac{1}{2}$  vs  $H_a : p \neq \frac{1}{2}$

Rejection rule: Reject  $H_0$  if  $X \leq 7$  or  $X \geq 18$

`qbinom(0.025, 25, 0.5) = 8`

`pbinom(8, 25, 0.5) = 0.054`

`pbinom(7, 25, 0.5) = 0.022`

Significance level:

`pbinom(7, 25, 0.5) + 1-pbinom(17, 25, 0.5) = 0.043`

Since we observed  $X = 17$ , we **fail to reject**  $H_0$

**P-value** =  $2 * (1 - \text{pbinom}(16, 25, 0.5)) = 0.11$

## Confidence interval for a proportion

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Suppose  $X \sim \text{binomial}(n=29, p)$  and we observe  $X = 24$ .

Consider the test of  $H_0 : p = p_0$  vs  $H_a : p \neq p_0$

**We reject  $H_0$  if**

$$\Pr(X \leq 24 \mid p = p_0) \leq \alpha/2 \quad \text{or} \quad \Pr(X \geq 24 \mid p = p_0) \leq \alpha/2$$

95% confidence interval for  $p$ :

The set of  $p_0$  for which a two-tailed test of  $H_0 : p = p_0$  would **not** be rejected, for the observed data, with  $\alpha = 0.05$ .

The “plausible” values of  $p$ .

# Example

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$X \sim \text{binomial}(n=29, p)$ ; observe  $X = 24$

Lower bound of 95% confidence interval:

Largest  $p_0$  such that  $\Pr(X \geq 24 \mid p = p_0) \leq 0.025$

Upper bound of 95% confidence interval:

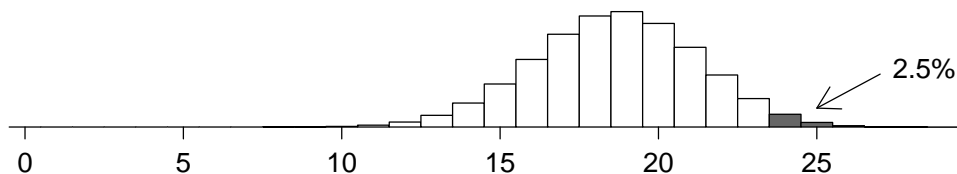
Smallest  $p_0$  such that  $\Pr(X \leq 24 \mid p = p_0) \leq 0.025$

In R: `binom.test(24, 29)`

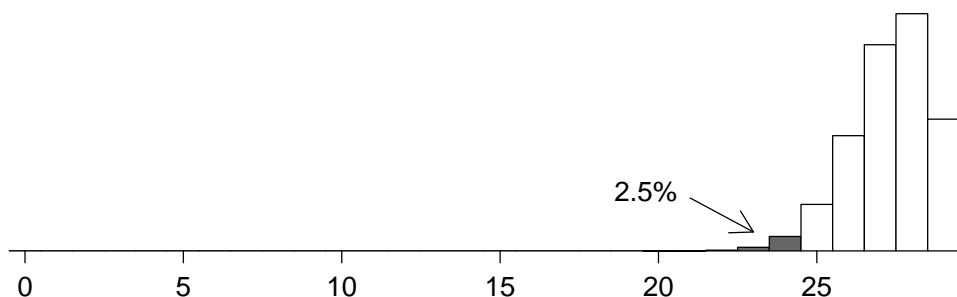
95% CI for  $p$ : (0.642, 0.942)

Note:  $\hat{p} = 24/29 = 0.83$  is not the midpoint of the CI

Binomial( $n=29, p=0.64$ )



Binomial( $n=29, p=0.94$ )



## Example 2

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$X \sim \text{binomial}(n=25, p)$ ; observe  $X = 17$

Lower bound of 95% confidence interval:

$p_L$  such that 17 is the 97.5 percentile of  $\text{binomial}(n=25, p_L)$

Upper bound of 95% confidence interval:

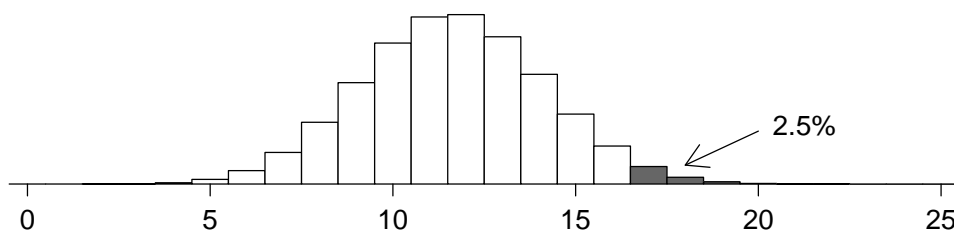
$p_H$  such that 17 is the 2.5 percentile of  $\text{binomial}(n=25, p_H)$

In R: `binom.test(17, 25)`

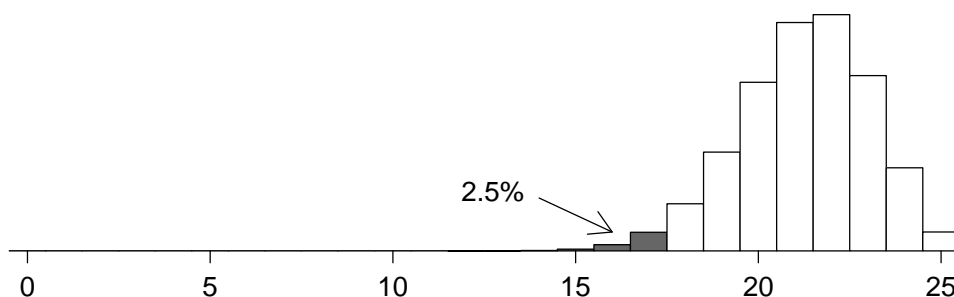
95% CI for  $p$ : (0.465, 0.851)

Again,  $\hat{p} = 17/25 = 0.68$  is not the midpoint of the CI

Binomial( $n=25, p=0.46$ )



Binomial( $n=25, p=0.85$ )



# The case $X = 0$

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Suppose  $X \sim \text{binomial}(n, p)$  and we observe  $X = 0$ .

Lower limit of 95% confidence interval for  $p$ : **0**

Upper limit of 95% confidence interval for  $p$ :

$p_H$  such that

$$\begin{aligned}\Pr(X \leq 0 \mid p = p_H) &= 0.025 \\ \implies \Pr(X = 0 \mid p = p_H) &= 0.025 \\ \implies (1 - p_H)^n &= 0.025 \\ \implies 1 - p_H &= \sqrt[n]{0.025} \\ \implies p_H &= 1 - \sqrt[n]{0.025}\end{aligned}$$

In the case  $n = 10$  and  $X = 0$ , the 95% CI for  $p$  is **(0, 0.31)**

## A mad cow example

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New York Times, Feb 3, 2004:

The department [of Agriculture] has not changed last year's plans to test 40,000 cows nationwide this year, out of 30 million slaughtered. Janet Riley, a spokeswoman for the American Meat Institute, which represents slaughterhouses, called that "**plenty sufficient from a statistical standpoint.**"

Suppose that the 40,000 cows tested are chosen **at random** from the population of 30 million cows, and suppose that **0** (or 1, or 2) are found to be infected.

How many of the 30 million total cows would we estimate to be infected?

What is the 95% confidence interval for the total number of infected cows?

| No. infected |       |            |
|--------------|-------|------------|
| Obs'd        | Est'd | 95% CI     |
| 0            | 0     | 0 – 2763   |
| 1            | 750   | 19 – 4173  |
| 2            | 1500  | 181 – 5411 |

## The case $X = n$

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Suppose  $X \sim \text{binomial}(n, p)$  and we observe  $X = n$ .

Upper limit of 95% confidence interval for  $p$ : **1**

Lower limit of 95% confidence interval for  $p$ :

$p_L$  such that

$$\begin{aligned}\Pr(X \geq n \mid p = p_L) &= 0.025 \\ \implies \Pr(X = n \mid p = p_L) &= 0.025 \\ \implies (p_L)^n &= 0.025 \\ \implies p_L &= \sqrt[n]{0.025}\end{aligned}$$

In the case  $n = 25$  and  $X = 25$ , the 95% CI for  $p$  is **(0.86, 1.00)**

## Large $n$ and medium $p$

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Suppose  $X \sim \text{binomial}(n, p)$ .

$$\begin{array}{lll} E(X) = np & SD(X) = \sqrt{np(1-p)} \\ \hat{p} = X/n & E(\hat{p}) = p & SD(\hat{p}) = \sqrt{\frac{p(1-p)}{n}} \end{array}$$

For large  $n$  and medium  $p$ ,  $\hat{p} \sim \text{normal}\left(p, \sqrt{\frac{p(1-p)}{n}}\right)$

Use 95% confidence interval  $\hat{p} \pm 1.96 \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$

**Unfortunately, this usually behaves poorly.**

Fortunately, you can just use `binom.test()`